

Search for B_s^0 - \overline{B}_s^0 oscillations in DELPHI using high- p_t leptons

The DELPHI Collaboration

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Abstract. Oscillations in the $B_s^0-\overline{B}_s^0$ system were studied in events selected from about 4.3 million hadronic Z^0 decays registered by DELPHI between 1992 and 2000.

This paper presents updates of two published analyses ([11,12]). The first analysis, which utilizes leptons emitted with large momentum transverse to a jet, was improved by means of a better algorithm for the vertex reconstruction and a new algorithm for flavour-tagging at production time. The second analysis, which utilizes D_s -lepton events, was improved by optimizing the treatment of proper time resolution.

No signal of B_s^0 oscillations was observed and limits on the mass difference between the physical B_s^0 states were obtained to be:

$$\begin{aligned} \Delta m_s &> 8.0 \text{ ps}^{-1} \text{ at the 95\% C.L.} \\ &\text{with a sensitivity of } \Delta m_s = 9.1 \text{ ps}^{-1} \end{aligned}$$

in the high p_t lepton analysis and

$$\begin{aligned} \Delta m_s &> 4.9 \text{ ps}^{-1} \text{ at the 95\% C.L.} \\ &\text{with a sensitivity of } \Delta m_s = 8.6 \text{ ps}^{-1} \end{aligned}$$

in the D_s -lepton analysis.

Previously published results on these analyses are superseded.

The combination of these results with those obtained in other independent analyses previously performed in DELPHI (D_s -hadron, exclusive B_s^0 , inclusive vertex) gives:

$$\begin{aligned} \Delta m_s &> 8.5 \text{ ps}^{-1} \text{ at the 95\% C.L.} \\ &\text{with a sensitivity of } \Delta m_s = 12.0 \text{ ps}^{-1}. \end{aligned}$$

1 Introduction

One of the main interests in B physics is the precise determination of the values of the ρ and η parameters, the two least known parameters of the CKM matrix [1] in the Wolfenstein parameterisation. The unitarity of this matrix can be visualized as a triangle in the $\rho - \eta$ plane. Several quantities which depend on ρ and η can be measured and, if the Standard Model is correct, they must yield, within measurement errors and theoretical uncertainties, compatible values for the these two parameters, inside measurement errors and theoretical uncertainties. One of these quantities is the mass difference (Δm_q) between CP eigenstates in the neutral B-meson systems $B_{d(s)}^0$.

In the Standard Model, $B_q^0 - \overline{B}_q^0$ ($q = d, s$) mixing is a direct consequence of second-order weak interactions. Starting with a B_q^0 meson produced at time $t=0$, the probability, \mathcal{P} , to observe a B_q^0 decaying at the proper time t can be written, neglecting effects from CP violation:

$$\mathcal{P}(B_q^0 \rightarrow B_q^0) = \frac{\Gamma_q}{2} e^{-\Gamma_q t} [\cosh(\frac{\Delta\Gamma_q}{2} t) + \cos(\Delta m_q t)].$$

Here $\Gamma_q = \frac{\Gamma_q^H + \Gamma_q^L}{2}$, $\Delta\Gamma_q = \Gamma_q^H - \Gamma_q^L$ and $\Delta m_q = m_q^L - m_q^H$, where L and H denote respectively the light and heavy physical states, $m_q^{L(H)}$ and $\Gamma_q^{L(H)}$ are the mass and total decay width of these states. The oscillation period gives a direct measurement of the mass difference between the two physical states. The Standard Model predicts that $\Delta\Gamma \ll \Delta m$ [1]. Neglecting a possible difference between

the B_s^0 lifetimes of the heavy and light mass eigenstates, the above expression simplifies to:

$$\mathcal{P}_{B_q^0}^{\text{unmix.}} = \mathcal{P}(B_q^0 \rightarrow B_q^0) = \frac{1}{2\tau_q} e^{-\frac{t}{\tau_q}} [1 + \cos(\Delta m_q t)]$$

and similarly:

$$\mathcal{P}_{B_q^0}^{\text{mix.}} = \mathcal{P}(B_q^0 \rightarrow \overline{B}_q^0) = \frac{1}{2\tau_q} e^{-\frac{t}{\tau_q}} [1 - \cos(\Delta m_q t)].$$

In the Standard Model, the $B_q^0 - \overline{B}_q^0$ ($q = d, s$) mixing frequency Δm_q (keeping only the dominant top quark contribution) can be expressed as follows:

$$\Delta m_q = \frac{G_F^2}{6\pi^2} |V_{tb}|^2 |V_{tq}|^2 m_t^2 m_{B_q} f_{B_q}^2 B_{B_q} \eta_B F\left(\frac{m_t^2}{m_W^2}\right). \quad (1)$$

In this expression G_F is the Fermi coupling constant; $F(x_t)$, with $x_t = \frac{m_t^2}{m_W^2}$, results from the evaluation of the box diagram and has a smooth dependence on x_t ; η_B is a QCD correction factor obtained at next-to-leading order in perturbative QCD. The dominant uncertainties in (1) come from the evaluation of the B-meson decay constant f_{B_q} and of the ‘‘bag’’ parameter B_{B_q} .

The two elements of the V_{CKM} matrix are equal to:

$$|V_{td}| = \lambda^3 \sqrt{(1-\rho)^2 + \eta^2}; \quad |V_{ts}| = \lambda^2, \quad (2)$$

neglecting terms of order $\mathcal{O}(\lambda^4)$. $|V_{ts}|$ is independent of ρ and η and is equal to $|V_{cb}|$.

Δm_d has been measured precisely by several experiments. Nevertheless this precision cannot be fully exploited due to the large uncertainty which originates in the evaluation of the non-perturbative QCD parameters. The ratio between the Standard Model expectations for Δm_d and Δm_s is given by the following expression:

$$\frac{\Delta m_d}{\Delta m_s} = \frac{m_{B_d} f_{B_d}^2 B_{B_d} \eta_{B_d}}{m_{B_s} f_{B_s}^2 B_{B_s} \eta_{B_s}} \frac{|V_{td}|^2}{|V_{ts}|^2}. \quad (3)$$

A measurement of the ratio $\frac{\Delta m_d}{\Delta m_s}$ gives the same type of constraint in the ρ - η plane, as a measurement of Δm_d , and this ratio only depends on f_{B_d}/f_{B_s} and B_{B_d}/B_{B_s} which can be predicted with better precision than the absolute values.

Using existing measurements which constrain ρ and η , except those on Δm_s , the distribution for the expected values of Δm_s can be obtained. It has been shown that Δm_s has to lie, at the 68% C.L., between 10.9 ps^{-1} and 18.1 ps^{-1} and is expected to be smaller than 21.5 ps^{-1} at the 95% C.L. [2].

2 The inclusive lepton analysis

For a description of the DELPHI detector and its performance, the reader is referred to [3]. The analysis described in this paper used precise tracking based on the silicon microvertex detector (VD) to reconstruct the primary and secondary vertices. To estimate the B-meson momentum and direction, the neutral particles detected in the electromagnetic and hadronic calorimeter and the reconstructed charged particle tracks were used. Muon identification was based on the hits in the muon chambers being associated with a track. Electrons were identified using tracks associated with a shower in the electromagnetic calorimeter. The dE/dx energy loss measurement in the Time Projection Chamber (TPC) and the Cherenkov light detected in the Ring Imaging Cherenkov detector (RICH) were used to separate pions (and also electrons or muons) from kaons and protons.

Charged particles were selected if they satisfied the following criteria: a momentum larger than $200 \text{ MeV}/c$, a track length larger than 30 cm , a relative momentum error smaller than 130%, a polar angle (θ , defined with respect to the beam axis) between 20° and 160° and an impact parameter with respect to the primary vertex, determined on an event-by-event basis, smaller than 4 cm in the xy plane (perpendicular to the beam axis) and 10 cm in z (along the beam direction). Neutral particles had to deposit at least 500 MeV in the calorimeters and their polar angles had to lie between 2° and 178° .

To select hadronic events it was required that more than 7 charged particles were accepted with a total energy larger than 15 GeV . The sphericity direction was determined using charged and neutral particles and its polar angle was required to satisfy $|\cos \theta_{\text{sphericity}}| < 0.8$. The event was divided into two hemispheres by a plane perpendicular to the sphericity axis. In each hemisphere the total energy from charged and neutral particles had to be

larger than 5 GeV . A total of about 4 million hadronic Z decays were selected from which 3.5 million were taken in the LEP I phase (1992–1995) and 0.5 million were collected as calibration data in the LEP II phase (1996–2000). Two different detector periods were analysed separately: 1992–1993 and 1994–2000. In the 1992 and 1993 data the vertex detector measured only the $R\phi$ coordinate (R being defined as $\sqrt{x^2 + y^2}$ and ϕ is the azimuthal angle), while from 1994 to 2000 the z coordinate was also measured.

Using tracks with vertex detector information, the primary vertex was fitted using the average beam spot as a constraint [4].

Jets were reconstructed using charged and neutral particles by the LUCCLUS [5] jet algorithm with a transverse momentum selection cut d_{join} of $6 \text{ GeV}/c$.

Muons were identified by combining the measured position of the muon chamber hits with the tracking information. The tracks of charged particles were extrapolated to the muon chambers and then associated and fitted to the hits. The muon identification algorithm is described in [3]. “Loose” identified muons with momenta larger than $3 \text{ GeV}/c$ were accepted, as well as “standard” and “tight” identified muons with momenta larger than $2 \text{ GeV}/c$.

The electron candidates were identified by combining the electromagnetic shower information from the High density Projection Chamber (HPC) with the particle ionization loss, dE/dx , measured by the TPC. A sizeable fraction of electrons originates from photon conversions. They were partially rejected if two oppositely charged particles formed a secondary vertex where the invariant mass was zero within measurement errors. The different variables were combined using a neural network [3]. Selection cuts on the output of the network were adjusted to vary with the particle momentum ensuring a constant efficiency between $2 \text{ GeV}/c$ and $30 \text{ GeV}/c$. Two levels of selection (“standard” and “tight”) were tuned to provide 75% and 65% efficiency for electrons from B semileptonic decays, respectively.

Samples of hadronic Z decays (4 million events) and of Z bosons decaying only into $b\bar{b}$ quark pairs (2 million events) were simulated using the Monte Carlo generator JETSET 7.3 [5] with DELPHI-tuned JETSET parameters and updated b and c decay Tables [6]. The detailed response of the DELPHI detector was simulated [3].

The principle of the oscillation measurements is as follows. Each of the charged and neutral particles measured in the event is assigned to one of the two hemispheres defined by the plane transverse to the sphericity axis. A “production tag”, correlated to the b/\bar{b} sign of the initial quark at the production point, is defined using both hemispheres; in the hemisphere containing the lepton, the decay time of the B hadron is evaluated and a “decay tag” is defined, correlated with the B/\bar{B} nature of the decaying hadron. The lepton charge defines the “decay tag”.

2.1 Discriminant analysis

Several algorithms and selections in this paper are based on discriminant variables. The method used to build a discriminant variable is briefly described here. Given two

classes (A and B) of events and N variables that are able to distinguish, with different efficiencies, between events from these two classes, the discriminant variable R is defined as:

$$R = \prod_{i=1,N} f^A(x_i)/f^B(x_i)$$

where $f^A(x_i)$ and $f^B(x_i)$ are the probability density functions for the variable i in the samples A and B , respectively (computed from simulated events).

R is often rescaled, for practical reasons, in a finite range: in $[-1,1]$ for flavour tagging ($X = (R - 1)/(R + 1)$) or in $[0,1]$ ($X = R/(R + 1)$) in all the other cases.

2.2 Measurement of the B decay proper time

The B decay proper time, t_B , is measured from the estimates of the B decay distance, d_B , and momentum, p_B :

$$t_B = \frac{d_B m_B}{p_B}. \quad (4)$$

2.2.1 B decay distance

The B decay distance has been obtained from the measurement of the distance between the secondary and the primary vertices, projected along the jet direction in the $x - y$ plane and evaluated along the jet direction.

All charged particles with an impact parameter smaller than 2 mm relative to the beam interaction point in the plane transverse to the beam direction have been used to reconstruct the primary vertex. The position of the beam spot has been used as a constraint. If the primary vertex fit had a χ^2 -probability smaller than 10^{-3} , an iterative procedure was applied which removed the track contributing most to the χ^2 at each iteration. In a simulated $b\bar{b}$ sample, this procedure has been found to reconstruct vertices with an accuracy of 80 μm in the horizontal x direction, where the beam spot has the larger spread, and 40 μm in the vertical y direction.

The B decay distance was evaluated using two different algorithms.

- Mini-jets algorithm.

The position of the B decay vertex and the B momentum are obtained by using an algorithm especially conceived for B hadron semileptonic decays. This algorithm is based on a classification of the charged particles in the event as B decay products or as emitted from the beam interaction point. The B secondary vertex is obtained by intersecting the trajectories of the lepton and of a D candidate. The lepton track and at least one of the charged particles assigned as a D decay product have to be associated to hits in the VD. Particles from fragmentation and from B decay products are all present in the jet which contains the lepton, so an approach has been developed to distinguish between them. Ignoring the lepton, charged particles belonging to the

jet are gathered into low-mass clusters, using LUCCLUS with d_{join} reduced to 0.5 GeV/ c and assuming that the particles are pions. Inside each cluster, the particles are ordered in decreasing values of their pseudo-rapidity relative to the cluster direction. Those having the largest pseudo-rapidity values and a momentum larger than 500 MeV/ c are then kept until the mass of the resulting system exceeds 2.2 GeV/ c^2 . Clusters which make an angle larger than 500 mrad relative to the jet direction are discarded. If a cluster contains more than one particle measured in the VD, a secondary vertex is obtained from the particles belonging to the cluster, a pseudo-D track candidate is constructed and the intersection of the pseudo-D track with the lepton trajectory is evaluated. If a cluster contains only one particle measured in the VD, its intersection with the lepton trajectory is evaluated. Among all these secondary vertices, the one which has the largest significance is kept¹. Having selected the cluster which contains a majority of D decay products and to reduce possible biases induced by this selection on the measured decay length of the B hadron, this cluster is used simply as a seed to find the other particles emitted by the D, which may be classified in other clusters. For this purpose, all particles present in the jet, including neutrals but not the lepton, are ordered by decreasing values of their pseudo-rapidity relative to the direction of the momentum sum of the previously retained particles. Particles are then added to the previously retained ones until the mass of the system exceeds 2.2 GeV/ c^2 . A new evaluation of the D candidate trajectory is then obtained and a secondary vertex is constructed with the lepton track. All of the retained particles are then called B decay products. Once the set of tracks which contains a majority of D decay products is selected, the B secondary vertex is obtained by intersecting the trajectories of the lepton and of the D candidate.

- Grid algorithm.

The idea is based on constructing a 3-dimensional grid, where the points represent all possible secondary vertices. This grid is built up around the B direction. The B direction is defined by the sphericity axis for two-jet events or by the axis of the jet closest to the lepton candidate. The resolution on the B direction is around 50 mrad in θ and 60 mrad in ϕ . All the charged-particle tracks, other than the lepton, are then assigned to the primary or to any possible secondary vertex (any point of the grid) according to a probability which depends on the impact parameters in the $R\text{-}\phi$ plane and along z (computed with respect to this candidate vertex), the momentum and the rapidity of the particle with respect to the jet axis. The particles are clustered as tracks belonging to the primary vertex and to the candidate secondary vertex. The overall probability of each

¹ The significance is defined as the distance between the primary and the secondary vertices along the jet direction, evaluated in the plane transverse to the beam axis, divided by its measurement error.

configuration is then computed. The best configuration is the one with the maximum probability and it defines primary and secondary tracks. The D vertex is computed using the tracks that have been assigned to the secondary vertex. A B vertex is then obtained by fitting the trajectories of the lepton, the D candidate and the B direction to a common vertex.

If this procedure fails, the evaluation of the B secondary vertex is performed by intersecting the B direction with the lepton track.

A discriminant analysis is then performed to choose the best B secondary vertex, trying to separate the secondary vertices having a resolution better or worse than $250 \mu\text{m}$. For each algorithm a discriminant variable is constructed (following the procedure explained in 2.1) using the error on the reconstructed decay length for the B and the D candidate (if any), the χ^2 of the B and the D (if any) secondary vertex as well as the number of VD hits for the tracks forming the B vertex. Vertices characterized by a high (low) value of this discriminant variable have a good (poor) resolution.

The algorithm with the largest value of the discriminant variable is chosen for the given hemisphere.

In order to clean up the sample from vertex candidates with poor resolution, hemispheres with a value of the discriminant variable smaller than 0.4 are rejected. The selection cut has been chosen, studying simulated events, as the highest value that allows the sensitivity of the Δm_s analysis to be kept almost unchanged.

The remaining sample is divided into five classes according to the value of the discriminant variable.

2.2.2 B momentum

The B momentum is also evaluated using two different algorithms.

In the first algorithm, the B momentum is determined in several steps using the mini-jets method. First, each event is divided into two hemispheres separated by the plane transverse to the sphericity axis which contains the beam interaction point. Then the four-momentum of the B-meson, P_{meas}^B , is evaluated by subtracting from the four-momentum of the hemisphere the sum of the four-momenta of the particles not selected as B decay products. Then, to have a better estimate of the B momentum, the measured energies and momenta are rescaled by a common factor (α) and a missing four-momentum corresponding to a zero mass particle is added (P_ν, \vec{P}_ν). Energy and momentum conservation, applied to the complete event, determine these unknowns:

$$\alpha \left(\vec{P}_{\text{hem1}} + \vec{P}_{\text{hem2}} \right) + \vec{P}_\nu = \vec{0} \quad (5)$$

$$\alpha (E_{\text{hem1}} + E_{\text{hem2}}) + P_\nu = 2 E_{\text{beam}}. \quad (6)$$

The mean value of α , determined using simulation, is 1.13. If the direction of the missing momentum lies within 400 mrad of the direction of the D- ℓ system, its energy is attributed

to the B to account for the missing neutrino. A better approximation to the B momentum is then obtained using the simulation, by correcting for the average difference between the above estimator and the true B momentum, parameterised as a function of the reconstructed B momentum. Finally a global fit is applied to all the measured quantities: the primary and secondary vertex positions (6 variables), and the momentum vectors of the lepton and of the D and B-mesons (9 variables). Three constraints are applied:

- the direction given by the two vertices and the direction of the B momentum should be the same (two angular constraints),
- the mass of the B-meson should be equal to the nominal B_s^0 mass.

The second algorithm evaluated the B energy using the BSAURUS package [7]. A discriminant analysis is then performed to choose the best B energy, trying to discriminate the candidates having a relative energy resolution better than 8%.

The choice between the two algorithms is made using the same method as the one described for the vertex resolution.

In order to clean up the sample from B candidates with poor momentum resolution, hemispheres with a value of the discriminant variable smaller than 0.3 are rejected. This value has been chosen using the same procedure as the one described for the vertex resolution.

The remaining sample is divided into three classes according to the value of the discriminant variable.

2.2.3 Parameterisation of the resolution functions

To have a detailed description of the time resolution, the distribution of the difference between the generated (t_{sim}) and reconstructed (t_{rec}) B decay proper times $\mathcal{R}_B(t_{\text{sim}} - t_{\text{rec}})$ has been parameterised by using the sum of Gaussian distributions with widths that depend on the generated decay time and on the sign of $t_{\text{sim}} - t_{\text{rec}}$. For about 10% of the events, because of the presence of charged-particle tracks coming from the beam interaction point, the reconstructed vertex coincides with the event main vertex. A Gaussian distribution, centred on $t_{\text{rec}} = 0$, has been used to account for the proper time distribution of these events.

The accuracy of the measurement of the B decay proper time is degraded for cascade semileptonic decays ($b \rightarrow c \rightarrow \ell X$), since the parameterisation of the difference between the true and measured B momentum and the strategy for the reconstruction of the B decay point have been tuned assuming a direct semileptonic decay of a B hadron. This has been verified using the simulation and a different parameterisation $\mathcal{R}_{BC}(t_{\text{sim}} - t_{\text{rec}})$ has been obtained for this category of leptons.

The function used to parameterise the resolution is:

- for the central part: three asymmetric Gaussian distributions² with identical central values. The param-

² An asymmetric Gaussian distribution consists of two Gaussian distributions with the same central value and amplitude but with different widths in the region $x < \mu$ and $x > \mu$.

eterisation contains 10 parameters: 3 fractions $f_{i=1,3}$ ($f_1 > f_2 > f_3$), 1 central value ($x_1 = x_2 = x_3 = \mu$) and 6 widths σ_i, σ'_i .

$$G^{(x_i=\mu, \sigma_i, \sigma'_i)}(x) \equiv \frac{1}{\sqrt{2\pi}} \frac{2}{\sigma_i + \sigma'_i} e^{-\frac{(x-\mu)^2}{2\sigma_i^2}}, x \leq \mu$$

$$\equiv \frac{1}{\sqrt{2\pi}} \frac{2}{\sigma_i + \sigma'_i} e^{-\frac{(x-\mu)^2}{2\sigma_i'^2}}, x \geq \mu \quad (7)$$

- for the primary vertex, a Gaussian distribution corresponding to the fraction f_4 , of central value x_4 and width σ_4 .

The general expression for the time resolution is:

$$R(t_{\text{sim}} - t_{\text{rec}}, t_{\text{sim}}) = \sum_{i=1,4} f_i G^{(x_i(t_{\text{sim}}), \sigma_i(t_{\text{sim}}), r_i(t_{\text{sim}}))}(t_{\text{sim}} - t_{\text{rec}}) \quad (8)$$

where $r_i \equiv \frac{\sigma_i}{\sigma'_i}$. The fractions f_i are normalised in the following way:

$$f_4 = 1 - f_1 - f_2 - f_3. \quad (9)$$

This parameterisation contains 12 free parameters which depend on the generated time. The time dependence of these parameters is given in Table 1. The 29 quantities from a to m (see Table 1) are determined by fitting the $t_{\text{sim}} - t_{\text{rec}}$ distribution from the simulation as a function of t_{sim} . Finally the simulated time distribution for accepted

Table 1. Time dependence of the parameters entering into the parameterisation of the resolution function, where $t_{\text{exp}} \equiv 1 - \exp(-\frac{t}{\tau})$

$$\begin{aligned} f_{1,2,3} &= a_{1,2,3} + b_{1,2,3} \cdot t_{\text{exp}} + c_{1,2,3} \cdot t_{\text{exp}}^2 \\ f_4 &= 1 - f_1 - f_2 - f_3 \\ \mu &= d + e \cdot t_{\text{exp}} + f \cdot t_{\text{exp}}^2 + g \cdot t_{\text{exp}}^3 \\ x_1 = x_2 = x_3 &= \mu \\ x_4 &= h + i \cdot t \\ \sigma_{1,2,3,4} &= \sqrt{j_{1,2,3,4}^2 + k_{1,2,3,4}^2 \cdot t_{\text{exp}}^2} \\ r_{1,2,3} &= l_{1,2,3} + m_{1,2,3} \cdot t_{\text{exp}} \\ r_4 &= 1 \end{aligned}$$

events is compared with an exponential distribution corresponding to the generated lifetime and an acceptance function, $\mathcal{A}(t_{\text{sim}})$, is obtained. In practice \mathcal{A} is constant over the accepted range of decay times between 0 and 12 ps. For events originating from light and charm quark flavours, the expected time distributions, $\mathcal{P}_H(t)$ and $\mathcal{P}_C(t)$, have been obtained from the simulation. The distributions are parameterised using a combination of three exponentials convoluted with Gaussian distributions.

Finally the sample was uniformly divided into 15 sub-samples according to 5 classes in decay-length resolution and 3 classes in momentum resolution. Figures 1 and 2 show, as an example, the distributions of $(t_{\text{rec}} - t_{\text{sim}})$ in different intervals of t_{sim} as indicated in each histogram for the events which have respectively the worst and the best proper time resolution for the 1994–2000 data.

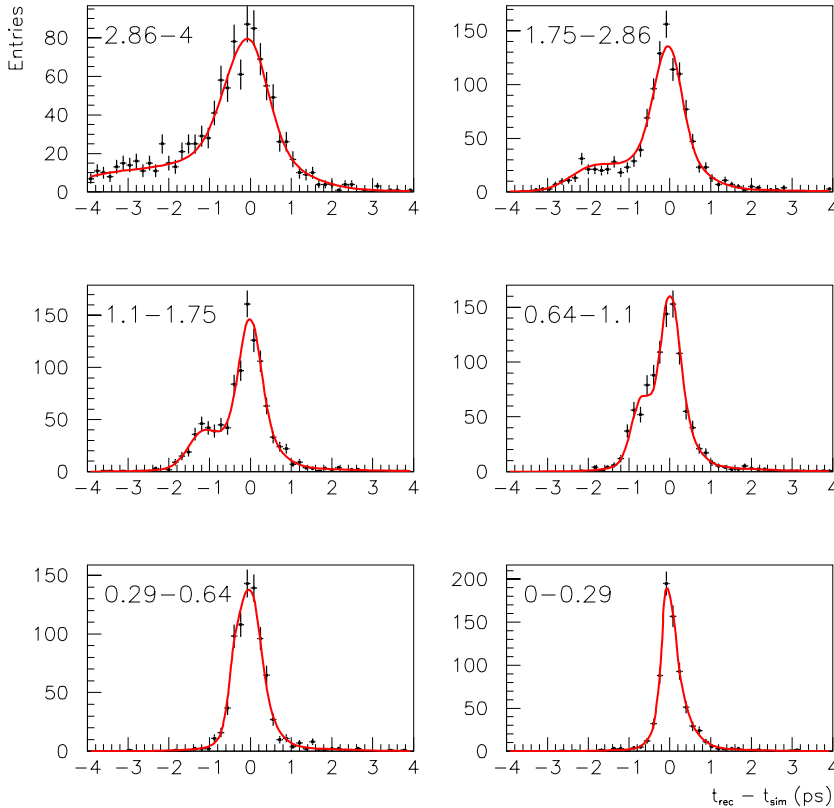


Fig. 1. The distributions of $t_{\text{rec}} - t_{\text{sim}}$ in different intervals of t_{sim} as indicated in top right corner of each plot (the unit is ps), for the worst class in proper-time resolution on simulated 1994–2000 data (points with error bars). The satellite peak appearing on the left part of the distribution corresponds to events in which the primary vertex has been reconstructed instead of the real B decay vertex. The curves correspond to the parameterisation explained in Sect. 2.2.3

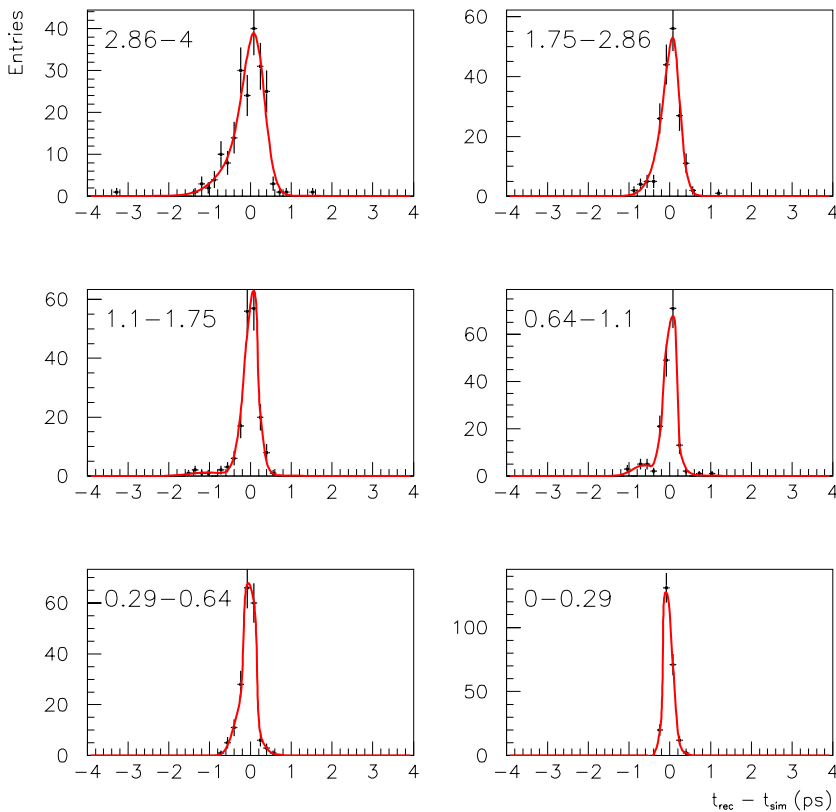


Fig. 2. Same as in Fig. 1 for the best class in proper-time resolution

2.3 Algorithm for tagging the b flavour at production time

The signature of the initial production of a b (\bar{b}) quark in the jet containing the B_s^0 or \overline{B}_s^0 candidate is determined in the opposite hemisphere, from a combination of different variables sensitive to the initial quark state. These variables are combined in a devoted neural network (BSAURUS [7]). The output of this neural network, $x_{\text{tag}}^{\text{opp}}$, varies between -1 and 1. Values of $x_{\text{tag}}^{\text{opp}}$ close to 1 (-1) correspond to a high probability that a given hemisphere contains a b (\bar{b}) quark in the initial state. This tagging is complemented in the same hemisphere with a simple discriminant variable using the information from tracks which are not attached to the reconstructed B vertex. Two variables are used:

- the mean jet charge which is defined as:

$$Q_{\text{hem}} = \frac{\sum_{i=1}^n q_i (\mathbf{p}_i \cdot \mathbf{e}_s)^\kappa}{\sum_{i=1}^n (\mathbf{p}_i \cdot \mathbf{e}_s)^\kappa}. \quad (10)$$

In this expression, q_i and \mathbf{p}_i are respectively the charge and the momentum of particles i , \mathbf{e}_s is the unit vector along the sphericity axis and $\kappa=0.6$;

- the rapidity of the charged particle with highest rapidity among the identified kaons in the same hemisphere (spectator kaon).

The output of the neural network in the opposite hemisphere, the mean jet charge and rapidity of the kaon in the same hemisphere are combined into a single discriminant variable x_{tag} as described in Sect. 2.1.

The distributions of $x_{\text{tag}}^{\text{opp}}$ and x_{tag} for simulation and data will be presented in Sects. 3.1 and 3.2.1.

2.4 Composition of the lepton sample

The muons were selected in the “tight” and “standard” categories while the electrons were required to belong to the “tight” category.

The efficiency to identify leptons and the hadronic contamination have been obtained by using the detailed simulation code of the DELPHI detector [3] and have been checked on data by using selected event samples such as $K_s^0 \rightarrow \pi^+\pi^-$, $Z^0 \rightarrow \mu^+\mu^-$, photons converted in front of the HPC, $\gamma\gamma \rightarrow \ell^+\ell^-$ and hadronic τ decays [3].

Candidate leptons are produced by semileptonic decays of B hadrons, D hadrons, and light mesons. They also can be misidentified hadrons or converted photons. Leptons from cascade decays have the wrong sign with respect to leptons from direct B decays for the identification of the b quark charge. Thus, selections have been defined on the lepton total, p , and transverse, p_t , momenta to minimize their contribution³. The fractions of the different categories of selected leptons, with p larger than 3 GeV/ c and p_t larger than 1.2 GeV/ c , have been measured using the real data by fitting the (p, p_t) distributions of the different components in the simulation to the corresponding distribution

³ These selections have been obtained by maximizing the product $(f_b^l - f_{bc}^l) \times \sqrt{N_l}$, where f_b^l and f_{bc}^l are the fractions of direct and cascade leptons in the total sample of N_l lepton candidates.

for lepton candidates observed in the data. The transverse momentum is computed with respect to the jet axis after having removed the lepton from the jet.

To study $B - \bar{B}$ oscillations, lepton candidates in the simulated event sample have been divided into four classes according to their sign relative to the sign of the heavy quark present in the decaying hadron:

- ℓ_b : lepton candidates, produced in B hadron decays, having the same sign as the b quark present inside the B hadron. This class contains leptons from direct b semileptonic decays including τ decay products and also, for example, those from cascade decays of the type: $\bar{B} \rightarrow \text{DDX}$ with $\bar{D} \rightarrow \ell^- X$. Misidentified hadrons with the same sign as the b quark are also included if they originate from a B hadron decay.
- ℓ_{bc} : when the candidate lepton has a sign opposite to the b quark charge.
- ℓ_c : when the candidate lepton is produced in a charmed hadron decay which is not coming from a B-meson decay.
- ℓ_h : for candidate leptons which are misidentified hadrons or leptons produced in light hadron decays or converted photons and which do not fall into the previous classes.

With these definitions, if a B hadron oscillates, lepton candidates in the ℓ_b and ℓ_{bc} classes change their sign, whereas those belonging to the other classes are not affected.

The semileptonic decay branching fractions used in the simulation have been modified to agree with the measured values. The fraction of fake lepton candidates has been verified in data with an accuracy of $\pm 10\%$ [8]. The Monte Carlo simulation is then corrected so that the rates of the physics processes and their simulation agree with present measurements. As an example, for 1994–2000 data, the fractions of lepton candidates in the worst and best classes in proper time resolution are:

$$\begin{aligned} f_b^l &= 67.6\%, f_{bc}^l = 15.4\%, f_c^l = 4.8\%, f_h^l = 12.1\% \\ &\text{(1994–2000 worst class)} \\ f_b^l &= 87.2\%, f_{bc}^l = 7.7\%, f_c^l = 3.2\%, f_h^l = 1.9\% \\ &\text{(1994–2000 best class)} \end{aligned} \quad (11)$$

Uncertainties attached to these fractions depend on those on the semileptonic branching fraction measurements, on the lepton identification procedure and on the rate of the fake lepton background. They will be discussed in Sect. 3.

To improve the separation between the class corresponding to the signal, ℓ_b , and the other classes, a set of six variables has been used:

- p , the momentum of the reconstructed lepton;
- p_t , the transverse momentum of the reconstructed lepton with respect to the jet axis after having removed the lepton from the jet;
- $y_{b\text{-tag}}$, the b tagging variable. This variable is defined as the probability for the track impact parameters relative to the beam interaction position, in the hemisphere opposite to the one containing the lepton, to be compatible

- with the one expected for light quark events [4]. Small values for this variable tag events containing b-hadrons;
- the flight distance between the B and the D mesons;
- the multiplicity at the secondary vertex;
- the lepton identification.

Two different discriminant variables were constructed (see Sect. 2.1). The first variable distinguishes between events in which the lepton candidate is coming from B decays (direct and cascade), D decays or is a misidentified hadron. The second one differentiates between leptons from direct and cascade B decays. The first discriminant variable uses all the above mentioned variables, while in constructing the other, the $y_{b\text{-tag}}$ and the lepton identification were removed. Two-dimensional distributions of the discriminant variables were then constructed for each of the proper-time resolution classes. Figure 3 shows the projections onto these discriminant variables. The fractions of the different components in the sample as a function of the two variables are also shown in Fig. 3. This information will be used on an event-by-event basis.

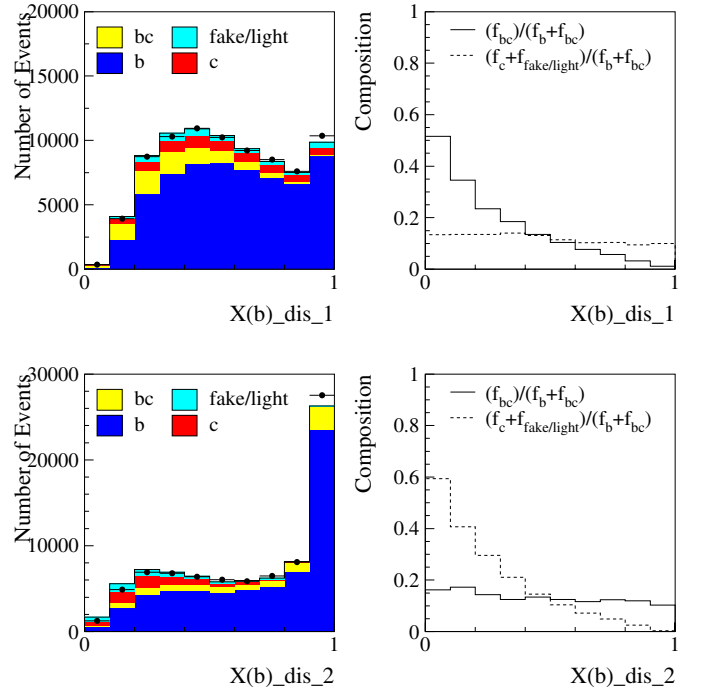


Fig. 3. The plots show the distribution of the discriminant variables used to distinguish between the events in which the lepton candidate is coming from direct B decays (bl), cascade decays (bcl), charm decays (c) and misidentified hadrons (fake). They are shown for 1994–2000 data (full dots). The plots on the top (bottom) show the projection onto the discriminant variable used to distinguish between events in which the leptons come from direct and cascade B decays (from B decays (bl or bcl) and the rest). The data/Monte Carlo agreement is shown on the left. The figures on the right show the composition of the simulated sample as a function of the value of the discriminating variable

Table 2. List of the values used for the most relevant parameters [1]

Parameter	Value
f_{B^+}, f_{B_d}	0.388 ± 0.013
f_{B_s}	0.106 ± 0.013
$f_{b\text{-baryons}}$	0.118 ± 0.020
τ_{B^+}	1.674 ± 0.018 ps
τ_{B_d}	1.542 ± 0.016 ps
τ_{B_s}	1.461 ± 0.057 ps
$\tau_{b\text{-baryons}}$	1.208 ± 0.051 ps
Δm_d	0.489 ± 0.008 ps $^{-1}$

2.5 The fraction of B hadrons of flavour q in the sample: f_{B_q}

Besides other inputs like product branching fractions involving characteristic signatures, precise information on f_{B_s} can be obtained by comparing the integrated oscillation rates of neutral B-mesons ($\overline{\chi}$), measured at LEP and SLD, and of \overline{B}_d^0 mesons only (χ_d), measured at LEP and at the $\Upsilon(4S)$:

$$\overline{\chi} = f_{B_d}\chi_d + f_{B_s}\chi_s \quad (12)$$

together with the normalisation condition and using isospin symmetry in the non-strange B-meson sector:

$$f_{B_d} + f_{B^+} + f_{B_s} + f_{b\text{-baryon}} = 1, \quad f_{B^+} = f_{B_d}. \quad (13)$$

Recent determinations of these fractions are reported in Table 2.

To increase the fraction of B_s^0 mesons in the sample, a set of four variables which are sensitive to the presence of strange B-mesons in a jet is defined:

- the number of kaons associated to the secondary vertex;
- the number of kaons coming from the primary vertex;
- the multiplicity of charged-particle tracks at the secondary vertex;
- the charge of the reconstructed secondary vertex.

Probability distributions of these variables have been combined to define two discriminant variables. Only the first three variables were used to discriminate between B_s^0 and the other neutral hadrons, while all the four variables were used to discriminate between B_s^0 and charged B-mesons. Two-dimensional distributions of the discriminant variables (see Sect. 2.1) were then constructed for each of the resolution classes. Figure 4 shows the projections onto the two discriminant variables for the 1994–2000 data, together with the variation of the B_s^0 fraction as a function of the two discriminating variables. This information will be used on an event-by-event basis.

2.6 Fitting procedure

Events have been classified on the basis of the charge of the lepton, Q_ℓ , and x_{tag} . They are considered as mixed if $x_{\text{tag}} \times Q_\ell < 0$ and as unmixed if $x_{\text{tag}} \times Q_\ell > 0$.

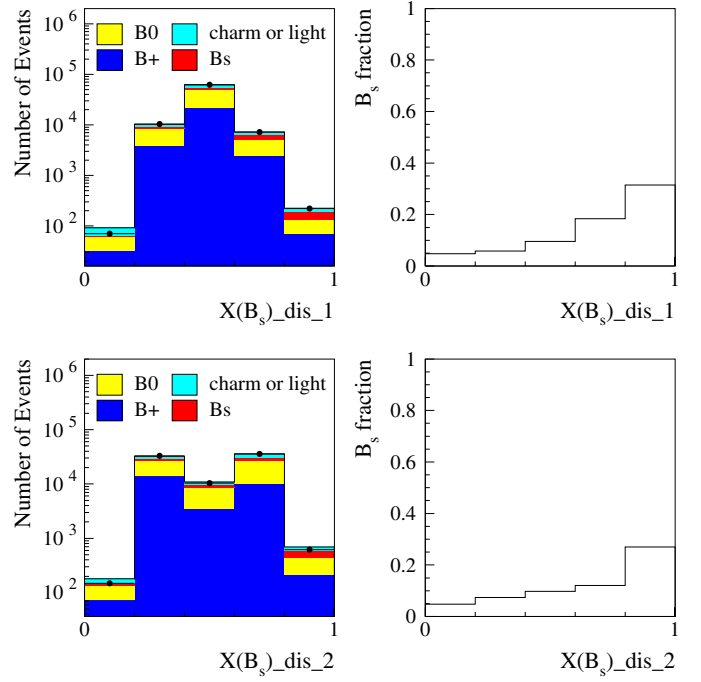


Fig. 4. The plots show the distribution of the discriminant variables used to distinguish between the events coming from B_s^0 decays from those coming from other neutral b-hadron decays (top) and from B^+ decays (bottom). They are shown for 1994–2000 data (full dots). On the left the data/Monte Carlo agreement can be appreciated, on the right the figures show the B_s^0 fraction as a function of the value of the discriminating variable

The number of events are 16448 (10160) and 26132 (15368) in the mixed and unmixed categories, respectively, for the 1994–2000 (1992–1993) data. So 68108 events have been used for the analysis.

In each event the probability to obtain, for a given measured proper time t , a like-sign pair (charge of the lepton equal to the charge of the quark at the production time: $\mathcal{P}^{\text{like}}(t)$) or an unlike-sign pair (charge of the lepton opposite to the charge of the quark at the production time: $\mathcal{P}^{\text{unlike}}(t)$) has been evaluated:

$$\begin{aligned}
\mathcal{P}^{\text{like}}(t) &= f_b^\ell \sum_q f_{B_q}^b \epsilon_b^{\text{tag}} (\mathcal{P}_{B_q}^{\text{mix.}}(t') \otimes \mathcal{R}_B(t' - t)) \\
&\quad + f_{bc}^\ell \sum_q f_{B_q}^{bc} (1 - \epsilon_b^{\text{tag}}) (\mathcal{P}_{B_q}^{\text{mix.}}(t') \otimes \mathcal{R}_{BC}(t' - t)) \\
&\quad + f_b^\ell \sum_q f_{B_q}^b (1 - \epsilon_b^{\text{tag}}) (\mathcal{P}_{B_q}^{\text{unmix.}}(t') \otimes \mathcal{R}_B(t' - t)) \\
&\quad + f_{bc}^\ell \sum_q f_{B_q}^{bc} \epsilon_b^{\text{tag}} (\mathcal{P}_{B_q}^{\text{unmix.}}(t') \otimes \mathcal{R}_{BC}(t' - t)) \\
&\quad + f_c^\ell \epsilon_c^{\text{like}} \mathcal{P}_C(t) + f_h^\ell \epsilon_h^{\text{like}} \mathcal{P}_H(t), \quad (14)
\end{aligned}$$

$$\begin{aligned}
 \mathcal{P}_{\text{unlike}}(t) &= f_b^\ell \sum_q f_{B_q}^b \epsilon_b^{\text{tag}} (\mathcal{P}_{B_q}^{\text{unmix.}}(t') \otimes \mathcal{R}_B(t' - t)) \\
 &+ f_{bc}^\ell \sum_q f_{B_q}^{bc} (1 - \epsilon_b^{\text{tag}}) (\mathcal{P}_{B_q}^{\text{unmix.}}(t') \otimes \mathcal{R}_{BC}(t' - t)) \\
 &+ f_b^\ell \sum_q f_{B_q}^b (1 - \epsilon_b^{\text{tag}}) (\mathcal{P}_{B_q}^{\text{mix.}}(t') \otimes \mathcal{R}_B(t' - t)) \\
 &+ f_{bc}^\ell \sum_q f_{B_q}^{bc} \epsilon_b^{\text{tag}} (\mathcal{P}_{B_q}^{\text{mix.}}(t') \otimes \mathcal{R}_{BC}(t' - t)) \\
 &+ f_c^\ell (1 - \epsilon_c^{\text{like}}) \mathcal{P}_C(t) + f_h^\ell (1 - \epsilon_h^{\text{like}}) \mathcal{P}_H(t). \quad (15)
 \end{aligned}$$

where the generated proper time is denoted t' ; f_b^ℓ , f_{bc}^ℓ , f_c^ℓ , f_h^ℓ are the fractions defined on an event-by-event basis as explained in Sect. 2.5; ϵ_b^{tag} is the tagging purity at production time (which is the fraction of events where the charge of the b at the production time is correctly assigned). In the fit procedure, the p.d.f. (probability density functions) to have the right (wrong) sign for the quark at production and decay times are directly used. To cope with possible differences between data and simulation, for the opposite hemisphere tag, the p.d.f. of $x_{\text{tag}}^{\text{oppo}}$ determined in the Δm_d fit (see Sect. 3.1) is used.

The p.d.f. for the same-hemisphere tagging variables are taken from simulated data.

The opposite and same-hemisphere tagging variables are combined in one single tagging variable x_{tag} .

In $Z^0 \rightarrow b\bar{b}$ events, $f_{B_q}^b$ and $f_{B_q}^{bc}$ are the fractions of direct and cascade decays, respectively, of B hadrons of flavour q in the sample. For direct decays it has been assumed that these fractions are the same as the corresponding production rates of the different B hadrons in b jets because very similar semileptonic partial widths are expected for all B hadrons (they are equal for \bar{B}_d^0 and B^- due to isospin invariance). The cascade decay fractions have been computed from Monte Carlo simulation after correcting the production rates.

In $Z^0 \rightarrow c\bar{c}$ events and for light flavours, ϵ_c^{like} and ϵ_h^{like} are the fractions of events classified as mixed candidates. Their values have been obtained using the p.d.f. from the simulation.

The functions $\mathcal{P}_{B_q}^{\text{mix.}}(t)$ and $\mathcal{P}_{B_q}^{\text{unmix.}}(t)$ have been given in the introduction for neutral B-mesons. For charged B-mesons and b -baryons, the decay time distribution has a simple exponential behaviour. These distributions have to be convoluted with the time resolution distributions $\mathcal{R}_B(t' - t)$ and $\mathcal{R}_{BC}(t' - t)$ for direct and cascade semileptonic B decays respectively, obtained from the simulation.

For $Z^0 \rightarrow c\bar{c}$ and $Z^0 \rightarrow$ light quark pair events, the reconstructed time distributions obtained in simulation have been fitted directly to provide $\mathcal{P}_C(t)$ and $\mathcal{P}_H(t)$.

The proper decay time distribution of real data is shown in Fig. 5 for the 1994–2000 sample with the result of the fit superimposed. The agreement obtained is compatible with the expected systematic uncertainty on the proper-time

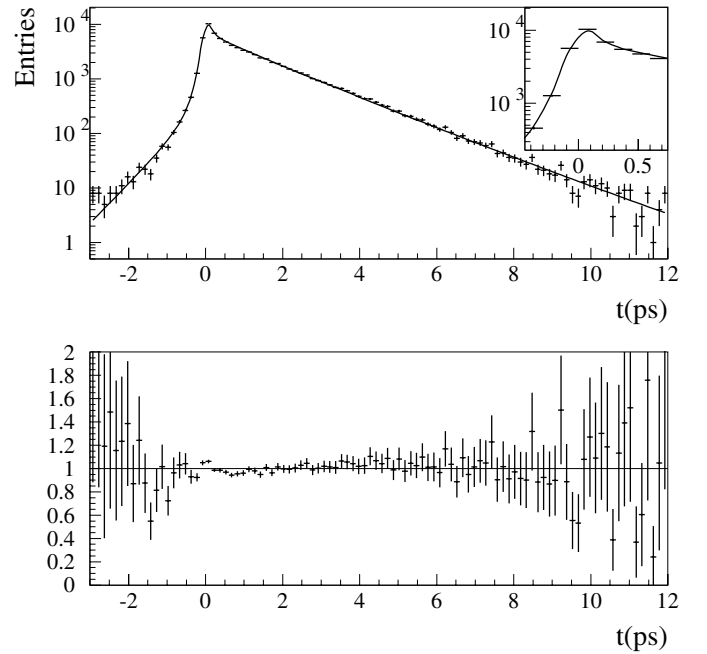


Fig. 5. The plot on the top shows the proper-time distribution in real data (1994–2000) with the fit described in the text superimposed. The bottom plot is obtained dividing, bin-by-bin, the distribution of the proper time measured in data by the value of the fitted function

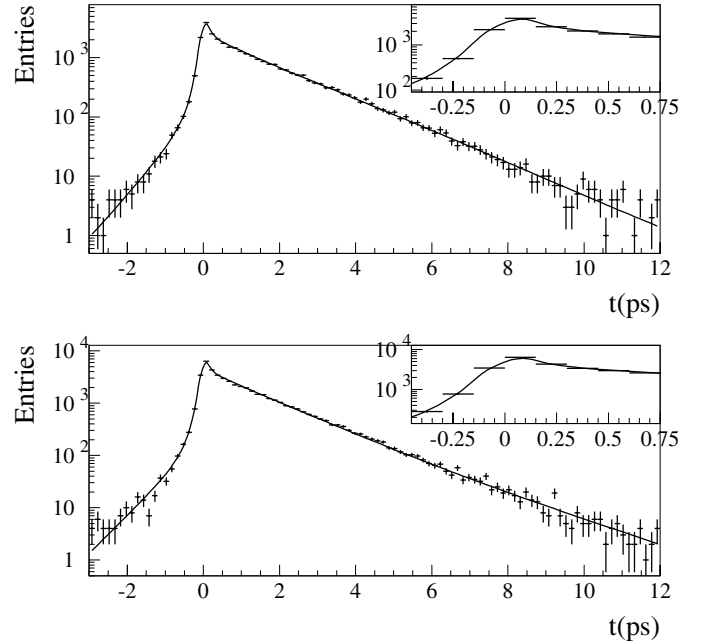


Fig. 6. The plots on the top (bottom) show the proper-time distributions in real data (1994–2000) with the fit superimposed for like-sign (unlike-sign) events

resolution. The same distribution for like and unlike-sign events separately is shown in Fig. 6.

The most relevant external inputs in the likelihood function are listed in Table 2.

3 Results of the inclusive lepton analysis

In this section the result of the measurement of the B_s^0 oscillation frequency, or a limit on its possible range, is given. The measurement of the B_d^0 oscillation frequency is also presented as a cross-check of the analysis technique.

3.1 Measurement of Δm_d

An unbinned maximum likelihood method has been applied to the set of classified events. The parameters corresponding to Δm_d and to the p.d.f. for the tagging in the opposite hemisphere in four different intervals of the $x_{\text{tag}}^{\text{opp}}$ discriminating variable have been fitted by minimizing the following function:

$$\mathcal{L} = - \sum_{\text{like-sign events}} \ln(\mathcal{P}^{\text{like}}(t)) - \sum_{\text{unlike-sign events}} \ln(\mathcal{P}^{\text{unlike}}(t)). \quad (16)$$

The results, converted to tagging purities, are given in Table 3. The final result is:

$$\Delta m_d = 0.456 \pm 0.021 \text{ ps}^{-1}.$$

The experimental distribution of the fraction of like-sign events versus the decay time is shown in Fig. 7. In the same figure the comparison between the fitted and simulated tagging p.d.f. is shown. The difference between these two p.d.f. shows that it is very important, where possible, to compute the distributions of interest on the data.

Table 3. Results from the fit on real data of Δm_d and of the tagging purity in four intervals of the discriminating variable as indicated in the last column. The fit was performed assuming that Δm_s is large ($\Delta m_s = 20 \text{ ps}^{-1}$ has been used)

variable	1992–1993	1994–2000	$x_{\text{tag}}^{\text{opp}}$ range
Δm_d (ps^{-1})	0.459 ± 0.036	0.455 ± 0.026	
ϵ_b^1	0.550 ± 0.008	0.550 ± 0.006	[0.00–0.25]
ϵ_b^2	0.644 ± 0.008	0.661 ± 0.006	[0.25–0.50]
ϵ_b^3	0.786 ± 0.008	0.762 ± 0.007	[0.50–0.75]
ϵ_b^4	0.873 ± 0.013	0.889 ± 0.008	[0.75–1.00]

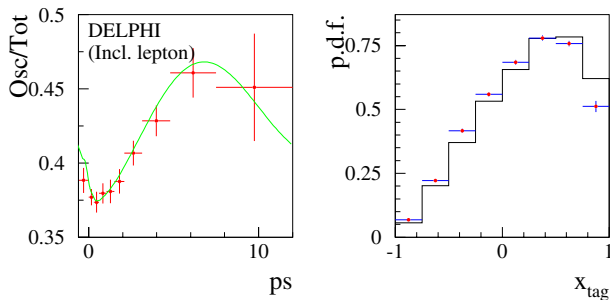


Fig. 7. Left plot: time dependence of the fraction of like-sign events. The curve is the result of the log-likelihood fit. The right histogram shows the probability density function (p.d.f.) for tagging a b quark from simulated data. The points with the error bars show the result of the fit to real data

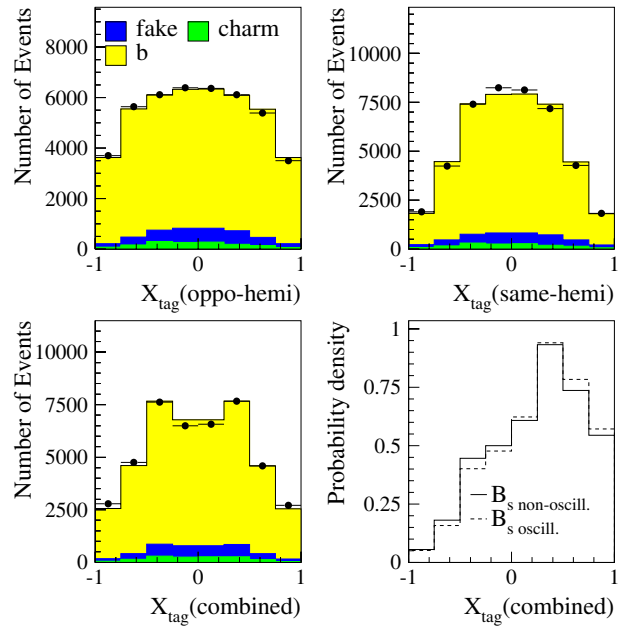


Fig. 8. The plots show the overall comparison between the distributions of the tagging variables in data and the p.d.f. used in the fit for opposite (top-left), same (top-right) hemisphere and the combined tagging (bottom-left) for the $B_s^0 - \overline{B}_s^0$ oscillation analysis. The plot on the bottom-right shows the expected p.d.f. for oscillating and non-oscillating B_s^0 , separately

The final tagging variable, x_{tag} , used in the Δm_s analysis is obtained by combining the tagging variable of the opposite hemisphere whose p.d.f. has been fitted on the data (fixing the Δm_d value to the world average) and the tagging variable of the same hemisphere whose p.d.f. has been computed from simulated data.

The overall agreement between the tagging p.d.f. used in the fit and the distributions of the tagging variable in the data is shown in Fig. 8, together with the tagging p.d.f. for oscillating and non-oscillating B_s^0 computed from simulated data. The distributions for oscillating and non-oscillating B_s^0 are expected to coincide only if the tagging in the same hemisphere is completely independent from the B_s^0 oscillation; the small difference found in the simulation has been taken into account in the fit by using different p.d.f. for the oscillation and non-oscillation hypothesis.

3.2 Limit on Δm_s

The limit on Δm_s has been obtained in the framework of the amplitude method. In this method [10], an oscillation amplitude, \mathcal{A} , is fitted for each assumed value of Δm_s . The equations for $\mathcal{P}_{B_s^0}^{\text{mix.}}$ and $\mathcal{P}_{B_s^0}^{\text{unmix.}}$ become:

$$\mathcal{P}_{B_s^0}^{\text{unmix.}} = \mathcal{P}(B_s^0 \rightarrow B_s^0) = \frac{1}{2\tau_s} e^{-\frac{t}{\tau_s}} [1 + \mathcal{A} \cos(\Delta m_s t)]$$

and

$$\mathcal{P}_{B_s^0}^{\text{mix.}} = \mathcal{P}(B_s^0 \rightarrow \overline{B}_s^0) = \frac{1}{2\tau_s} e^{-\frac{t}{\tau_s}} [1 - \mathcal{A} \cos(\Delta m_s t)].$$

For $\mathcal{A} = 1$, the standard time distribution expressions for mixed and unmixed candidates given in Sect. 1 are recovered. A measurement of the amplitude is obtained for each value of Δm_s . It has been verified that the log-likelihood distribution has a parabolic behaviour around its minimum.

In this approach, it is also easy to compute the probability of excluding a given value of Δm_s with the studied channel. It has to be assumed that the real value of Δm_s is very large and in practice not accessible with the present experimental sensitivity. The expected value of the amplitude is then equal to zero. All measured values of \mathcal{A} which satisfy $\mathcal{A} < 1 - 1.645\sigma_{\mathcal{A}}$ are such that the corresponding value of Δm_s is excluded at the 95% C.L. . . The Δm_s value satisfying $1.645 \sigma_{\mathcal{A}} = 1$ is then defined as the sensitivity of the analysis.

It is also possible, in this framework, to deduce the log-likelihood function $\Delta \log \mathcal{L}^\infty(\Delta m_s)$ referenced to its value obtained for $\Delta m_s = \infty$. The log-likelihood values can be easily deduced from \mathcal{A} and $\sigma_{\mathcal{A}}$ using the expressions given in [10]:

$$\begin{aligned} \Delta \log \mathcal{L}^\infty(\Delta m_s) &= \frac{1}{2} \left[\left(\frac{\mathcal{A} - 1}{\sigma_{\mathcal{A}}} \right)^2 - \left(\frac{\mathcal{A}}{\sigma_{\mathcal{A}}} \right)^2 \right] \\ &= \left(\frac{1}{2} - \mathcal{A} \right) \frac{1}{\sigma_{\mathcal{A}}^2}, \end{aligned} \quad (17)$$

$$\Delta \log \mathcal{L}^\infty(\Delta m_s)_{\text{mix}} = -\frac{1}{2} \frac{1}{\sigma_{\mathcal{A}}^2}, \quad (18)$$

$$\Delta \log \mathcal{L}^\infty(\Delta m_s)_{\text{nomix}} = \frac{1}{2} \frac{1}{\sigma_{\mathcal{A}}^2}. \quad (19)$$

The last two equations give the average log-likelihood value when $\Delta m_s = \Delta m_s^{\text{true}}$ (*mixing case*) and when Δm_s is different from Δm_s^{true} (*no-mixing case*).

Using the amplitude approach (Fig. 9), and considering only the statistical uncertainties, a 95% C.L limit on Δm_s was set:

$$\begin{aligned} \Delta m_s &> 8.0 \text{ ps}^{-1} \text{ at the 95\% C.L.} \\ &\text{with a sensitivity of } \Delta m_s = 9.6 \text{ ps}^{-1}. \end{aligned} \quad (20)$$

A check of the overall procedure has been performed using simulated data (Fig. 10) with different values of the B_s^0 purity. It can be seen that the minimum of the likelihood corresponds to the generated value of Δm_s and that its significance increases with the B_s^0 purity. This check gives some confidence that the Monte Carlo parameterisations are correct.

3.2.1 Systematic uncertainties on Δm_s

Systematics have been evaluated by varying, according to their respective uncertainties, the values of the parameters which were kept constant in the evaluation of the log-likelihood function. For each parameter, the variation of the amplitude and of its measurement error were taken into

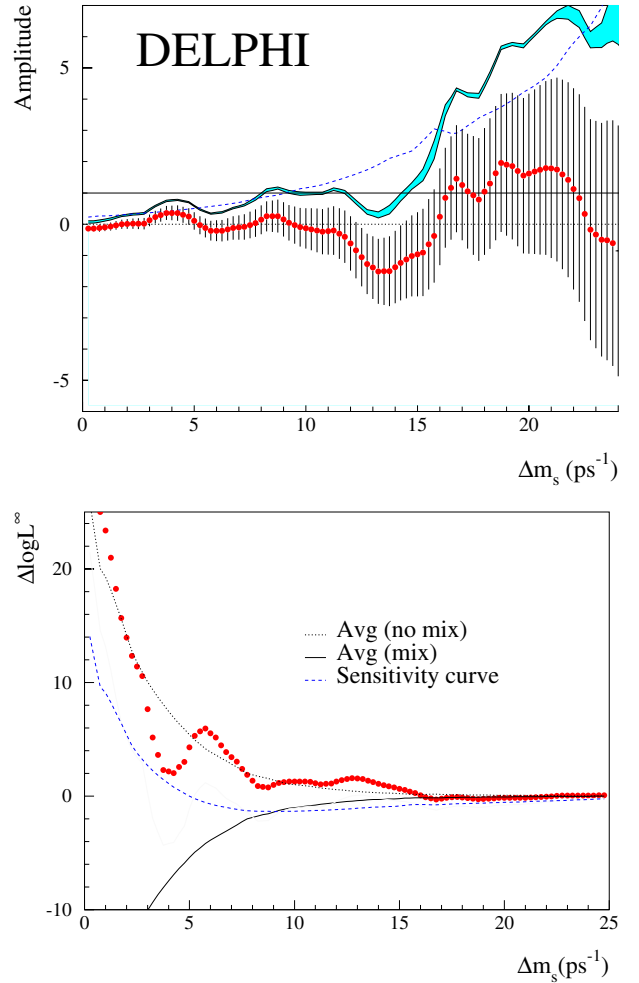


Fig. 9. DELPHI inclusive lepton analysis: the upper plot shows the variation of the oscillation amplitude \mathcal{A} as a function of Δm_s . The filled area shows the variation of the contour corresponding to $\mathcal{A} + 1.645\sigma_{\mathcal{A}}$ when the systematic uncertainty is included. The dotted line shows the sensitivity. The likelihood referenced to $\Delta m_s = \infty$ (lower plot), represented by points, has been deduced from the amplitude spectrum using the formula given in [10] (see Sect. 3.2)

account in the evaluation of the systematic uncertainty. This was done in the following way [10]:

$$\sigma_{\mathcal{A}}^{\text{sys}} = \mathcal{A} - \mathcal{A}_i + (1 - \mathcal{A}_i) \frac{\sigma_{\mathcal{A}}^{\text{stat}} - \sigma_{\mathcal{A}_i}^{\text{stat}}}{\sigma_{\mathcal{A}_i}^{\text{stat}}} \quad (21)$$

where \mathcal{A} , $\sigma_{\mathcal{A}}$ (\mathcal{A}_i , $\sigma_{\mathcal{A}_i}$) indicate the values and the errors on the amplitude after (before) the parameter variation; σ^{stat} is the statistical error on the amplitude. The following systematics have been considered:

- Δm_d , ϵ_b^i .

The systematic error due to Δm_d has been evaluated by varying the central value given in Table 2 by its standard deviation.

The tagging purity in the opposite hemisphere, ϵ_b^i , has been obtained from a fit to the real data (see Sect. 3.1 and Table 3). The systematics have been evaluated by

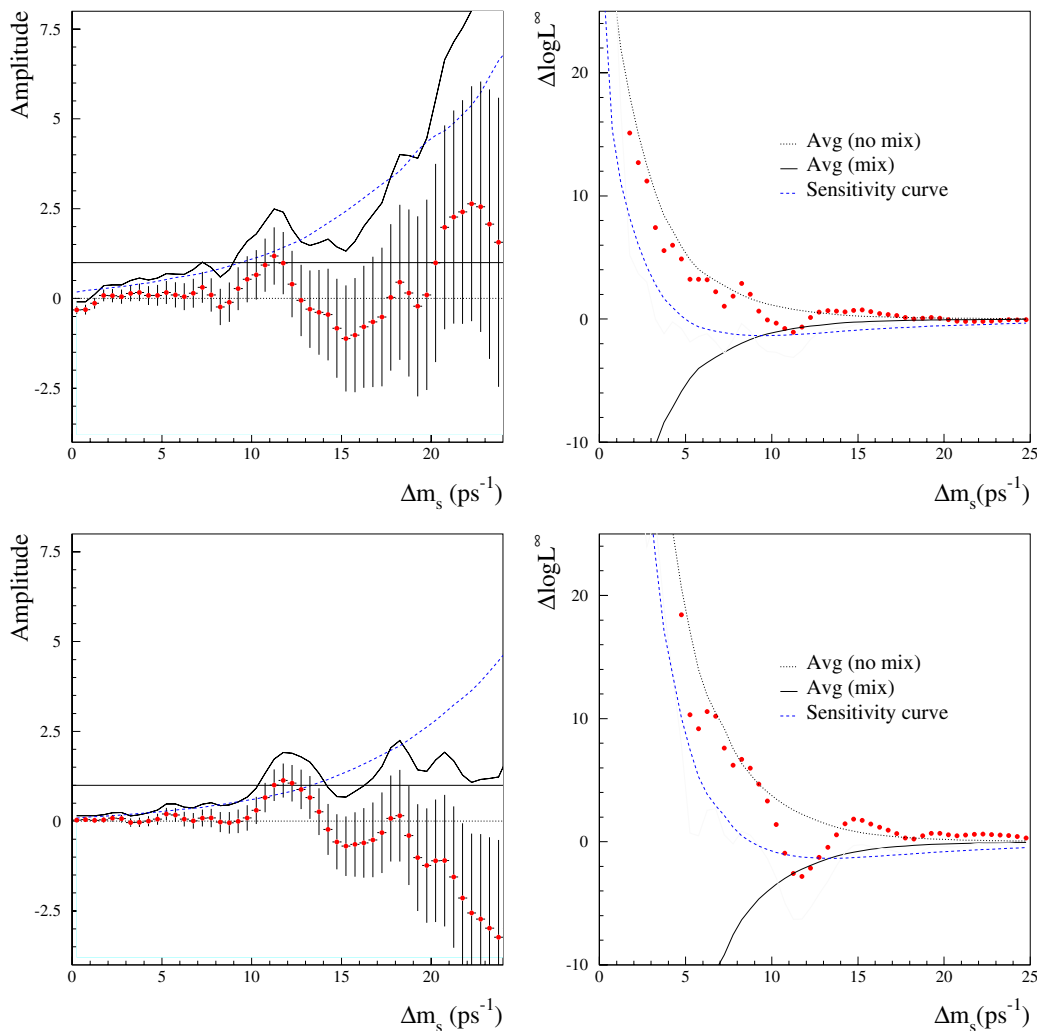


Fig. 10. Check of the overall fit procedure using simulated data. The plots on the top show the variation of the oscillation amplitude \mathcal{A} as a function of Δm_s (left) and the likelihood referenced to $\Delta m_s = \infty$ (right) for the simulated data with standard B_s^0 purity. The plots on the bottom give the same information if the B_s^0 purity is increased to 30%. The signal is expected (and seen) at 11.2 ps^{-1}

varying the shape of the variable of the opposite hemisphere tagging in accordance with the fitted errors in a way to maximize the data and Monte Carlo agreement in Fig. 8. The shape has been modified using a single parameter whose effect, on the mean purity, is ± 0.005 (± 0.007) in the 1992–1993 (1994–2000) data.

A conservative variation of the shape of the same hemisphere tagging corresponding to an integral variation of ± 0.010 has been also added.

– *Fractions of leptons.*

A relative variation of $\pm 20\%$ has been applied on the fraction of fake leptons, leptons from charm and cascade semileptonic decays. This variation is equal to the uncertainty on the measurement of the fake lepton rate in data. These changes in the parameters have been compensated by a corresponding variation of the fraction of direct leptons.

The systematics coming from the uncertainties on the shape of the discriminating variable distribution have been evaluated by forcing, with appropriate reweighting, an exact match between the distributions in data and simulation in Fig. 3.

– *Production rates of B hadrons.*

Following the procedure described in Sect. 2.5 the value

of f_{B_s} was varied inside its measured error.

The systematics coming from the uncertainties on the shape of the discriminating variable distribution has been evaluated by forcing, with appropriate reweighting, an exact match between the distributions in data and simulation in Fig. 4.

– *Systematics from the resolution of the B decay proper time.*

If the errors are Gaussian, the oscillation amplitude is damped by a factor ρ because of the finite accuracy in the decay time σ_t :

$$\rho = e^{-(\Delta m_s \sigma_t)^2 / 2} \quad (22)$$

where σ_t receives two contributions: from decay distance and from momentum measurements.

At small decay times, the accuracy on t depends mainly on the error on the decay distance. This quantity was measured using simulated events, after tuning the track reconstruction efficiencies and measurement errors to match the real data. For this purpose, charged particles emitted at angles smaller than 30° from the horizontal plane have been selected, in order to benefit from the precise definition of the beam position in the ver-

Table 4. List of the various contributions to the systematic error for different Δm_s values

Contribution to the systematic error	5 ps ⁻¹	10 ps ⁻¹	15 ps ⁻¹	20 ps ⁻¹
$f_{B_s^0}$	0.04	0.03	0.03	0.02
Fake lepton fraction	0.02	0.02	0.02	0.02
Lepton composition (p.d.f. shape)	0.01	0.02	0.05	0.04
B-hadron fractions (p.d.f. shape)	0.01	0.02	0.03	0.03
Tagging purity (p.d.f. shape)	0.01	0.01	0.02	0.03
Flight length resolution	0.02	0.07	0.31	0.33
Momentum resolution	0.01	0.04	0.12	0.13
Background proper-time shape	0.01	0.01	0.01	0.01
Subdivision in resolution classes	0.03	0.07	0.14	0.26
Total systematic error	0.14	0.28	0.68	0.72
Statistical error	0.30	0.65	1.35	2.56

tical direction. The details of the tuning procedure are described in [4]. After the tuning, the agreement between real and simulated data on the decay distance error has been evaluated from the width of the negative part of the flight distance distribution, for events which are depleted in b-hadrons. The difference of the widths in real data and in simulation is found to be approximately 10%.

The uncertainty on the momentum resolution was estimated to be $\pm 10\%$. This number was obtained by comparing the reconstructed B momentum in a hemisphere with the expected momentum in that hemisphere, obtained using energy and momentum conservation, for data and simulation; it was found that the momentum resolution agreed to better than $\pm 10\%$.

The systematic error coming from the uncertainties on the resolution functions is evaluated by varying by $\pm 10\%$ the two parameters describing the linear time dependence of the narrower Gaussian (see Table 1). A variation of $\pm 10\%$ of the resolution for the background events is also considered.

An important check has been carried out to verify that the subdivision in resolution classes does not introduce differences between the real and simulated data: the variation of the amplitude error as a function of Δm_s for real and simulated data has been compared. After rescaling data and Monte Carlo simulation to the same number of B_s^0 events, the ratio of the amplitude errors is compatible with 1 within $\pm 10\%$. An extra systematic uncertainty has been added by allowing a $\pm 10\%$ variation of the amplitude error.

The previous quantities have been varied separately. Table 4 summarizes the various contributions to the systematic error for different Δm_s values. The global systematic error is obtained by taking the quadratic sum of the individual systematics.

Including the present evaluation of systematics in the measured amplitude, the 95% C.L. limit is:

$$\Delta m_s > 8.0 \text{ ps}^{-1} \text{ at the 95\% C.L.} \quad (23)$$

with a sensitivity of $\Delta m_s = 9.1 \text{ ps}^{-1}$.

4 Update of the $D_s^\pm \ell^\mp$ analysis

$B_s^0 - \overline{B}_s^0$ oscillations have also been studied using an exclusively reconstructed D_s meson correlated with a lepton of opposite charge emitted in the same hemisphere:

$$\overline{B}_s^0 \longrightarrow D_s^+ \ell^- \overline{\nu}_\ell X.$$

Details of the analysis are described in [11]. A limit at the 95% C.L. on the mass difference between the physical B_s^0 states was obtained to be $\Delta m_s > 7.4 \text{ ps}^{-1}$, with a sensitivity of $\Delta m_s = 8.4 \text{ ps}^{-1}$. The weak point of this analysis was the rapid increase of the error on the amplitude as a function of Δm_s . It was due to a global parameterisation of the proper-time resolution. An improvement is obtained here by using the information of the proper-time resolution on an event-by-event basis. A discriminant analysis is performed to separate B secondary vertices with an expected good resolution (better than $250 \mu\text{m}$). This discriminant variable is constructed by using:

- the error on the reconstructed decay length for the B and the D candidate;
- the χ^2 of the B and the D secondary vertices;
- the D decay length divided by its error.

Another discriminant variable has been constructed to separate events having a momentum resolution better or worse than 8%. The variables used are:

- the reconstructed momentum;
- the D_s mass.

The two discriminant variables and their variation versus the error due to the decay length resolution, $\sigma_L = m_B/p_B \sigma(d_B)$, and relative momentum error, σ_{p_B}/p_B , are shown in Fig. 11.

The analysis is repeated for all the decay channels and separately for the 1992–1993 and 1994–1995 data. The proper-time resolution, $\sigma_t^2 = \sigma_L^2 + t^2(\sigma_p/p)^2$, is then obtained for each pair of values of the two discriminant variables and used in the fit procedure on an event-by-event basis.

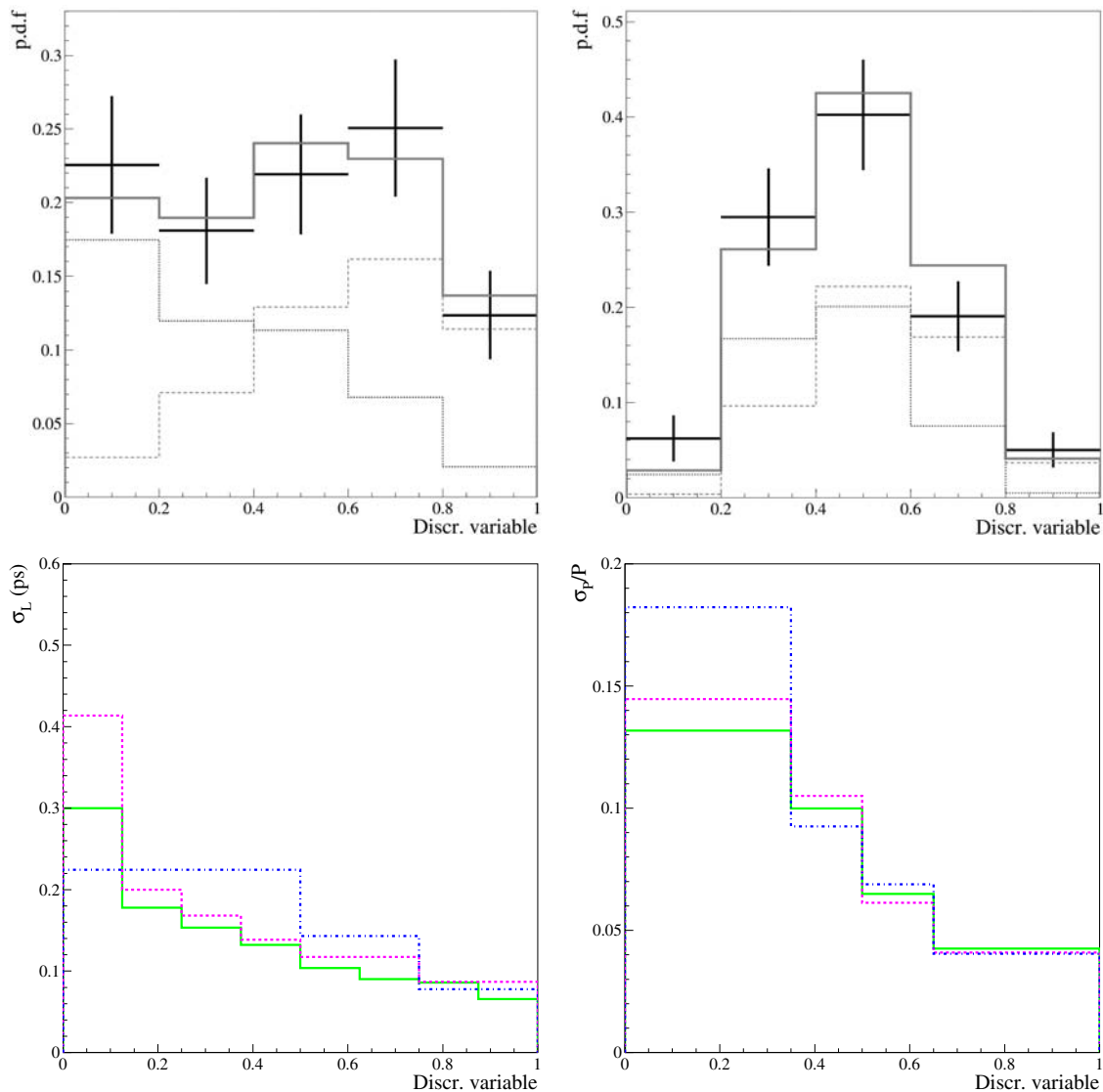


Fig. 11. Data/Monte Carlo comparison (represented with error bars and full line histogram respectively) for the discriminant variable for σ_L (upper left) (the events having a resolution better (worse) than $250 \mu\text{m}$ (dashed histograms) have a value nearer to 1 (0)) and for σ_p/p (upper right) (the events having a resolution better (worse) than 8% (dashed histograms) have a value nearer to 1 (0)). The lower plots show the evolution of σ_L and σ_p/p as a function of the discriminant variable. The continuous line is for the $\phi\pi$ and K^*K decay modes, the dot-dashed line is for the $\phi\ell\nu_\ell$ decay mode and the dashed line is for the K^0K decay mode. Only the width of the narrowest Gaussian is shown

A check has been done to verify, using simulated data, the new treatment of the proper-time resolution. The resolution obtained on an event-by-event basis by summing up all the contributions from the Gaussian distribution of widths σ_t reproduces well the overall resolution.

Another check has been performed by comparing the variation of the amplitude error, σ_A , as a function of Δm_s for real and simulated data. The agreement found indicates that the use of the proper-time resolution on an event-by-event basis does not introduce any significant difference in the error on the amplitude between real and simulated data. An extra systematic uncertainty is added by allowing a $\pm 10\%$ variation of the amplitude error.

Using the amplitude approach (Fig. 12) a 95% C.L limit on Δm_s was set:

$$\Delta m_s > 4.9 \text{ ps}^{-1} \text{ at the 95\% C.L.} \quad (24)$$

with a sensitivity of $\Delta m_s = 8.6 \text{ ps}^{-1}$.

The error on the amplitude at 20.0 ps^{-1} has been decreased by a factor 1.7 from that presented in [11] and also the sensitivity has improved.

5 Summary and combined limit on Δm_s

Using data registered with the DELPHI detector between 1992 and 2000 and considering the correlation between the sign of the charge of a lepton emitted at large transverse momentum relative to its jet axis, with the sign of a discriminating variable which uses several parameters of the

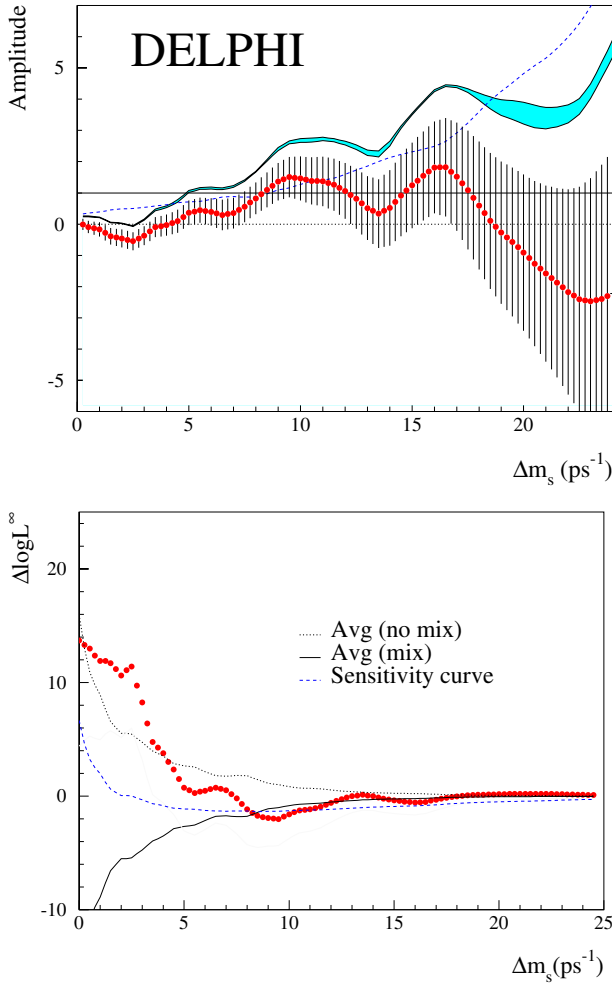


Fig. 12. DELPHI $D_s^\pm \ell^\mp$ analysis: the upper plot shows the variation of the oscillation amplitude \mathcal{A} as a function of Δm_s . The filled area shows the variation of the contour corresponding to $\mathcal{A} + 1.645\sigma_{\mathcal{A}}$ when the systematic uncertainty is included. The dotted line shows the sensitivity. The likelihood referenced to $\Delta m_s = \infty$ (lower plot), represented by points, has been deduced from the amplitude spectrum using the formula given in [10] (see Sect. 3.2)

event to define mixed and unmixed candidates, the value of the mass difference between B_d^0 mass eigenstates has been measured to be:

$$\Delta m_d = 0.456 \pm 0.021 \text{ ps}^{-1} \quad (25)$$

where the error accounts only for the statistical uncertainty. Using the amplitude approach, a 95% C.L. limit on Δm_s was set:

$$\Delta m_s > 8.0 \text{ ps}^{-1} \text{ at the 95\% C.L.} \quad (26)$$

with a sensitivity of $\Delta m_s = 9.1 \text{ ps}^{-1}$.

The previously published Δm_s analysis based on D_s -lepton events [11] has been updated improving the treatment of the proper-time resolution and the sensitivity. Using the amplitude approach, a 95% C.L. limit on Δm_s has

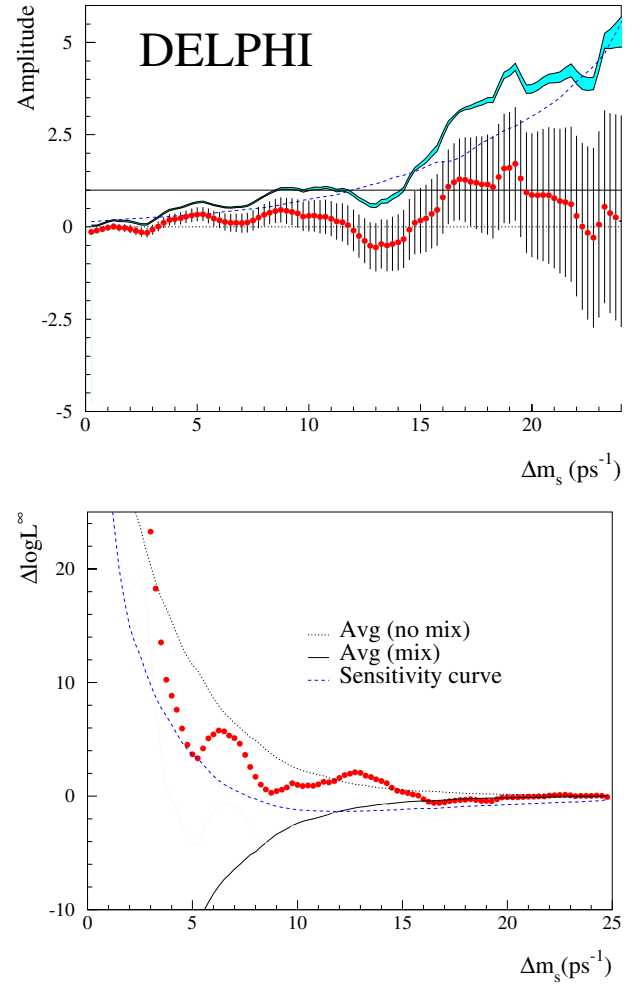


Fig. 13. DELPHI combined analysis: variation of the oscillation amplitude \mathcal{A} as a function of Δm_s . The filled area shows the variation of the contour corresponding to $\mathcal{A} + 1.645\sigma_{\mathcal{A}}$ when the systematic uncertainty is included. The dotted line shows the sensitivity. The likelihood referenced to $\Delta m_s = \infty$ (lower plot), represented with points, has been deduced from the amplitude spectrum using the formula given in [10] (see Sect. 3.2)

been set:

$$\Delta m_s > 4.9 \text{ ps}^{-1} \text{ at the 95\% C.L.} \quad (27)$$

with a sensitivity of $\Delta m_s = 8.6 \text{ ps}^{-1}$.

DELPHI has performed three other Δm_s analyses using exclusively reconstructed B_s^0 mesons [13], D_s -hadron events [13], and the inclusively reconstructed vertices [14]. These analyses on Δm_s have been combined, taking into account correlations between systematics in the different amplitude measurements (Fig. 13). The inclusive vertices analysis, the inclusive lepton analysis and the $D_s \ell$ analysis are statistically uncorrelated (high p_T leptons have been excluded from the vertex sample, $D_s \ell$ events have been excluded from the inclusive lepton sample); the small statistical correlation between the D_s -hadron events, the exclusively reconstructed B_s^0 mesons and the rest of the analyses has been neglected. The combined lower limit on

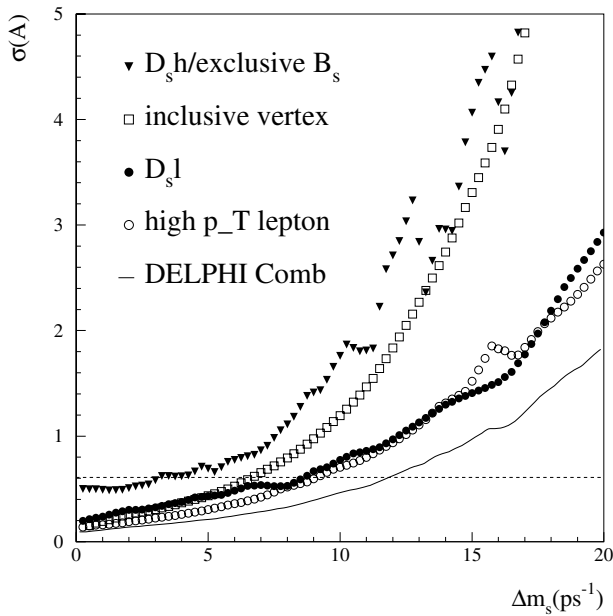


Fig. 14. Measured uncertainty on the amplitude as a function of Δm_s for all DELPHI analyses. For reference the line $\sigma = 1/1.645$ is also drawn: the abscissa of the intersection with each error curve is the sensitivity of the analysis

Δm_s turned out to be:

$$\Delta m_s > 8.5 \text{ ps}^{-1} \text{ at the 95\% C.L.} \quad (28)$$

with a sensitivity of $\Delta m_s = 12.0 \text{ ps}^{-1}$.

The sensitivity without systematics is of 12.5 ps^{-1} . The second excluded region between 11.8 ps^{-1} and 14.0 ps^{-1} (Fig. 13) is above the combined sensitivity and is not considered further.

The variation of the uncertainty on the amplitude as a function of Δm_s for all the analyses is given in Fig. 14.

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